

Some facts concerning ratios of random variables

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Concerning a question of Prof. T. Cacoullos, I pointed out the following:

If X and Y are (simply) identically distributed,

$$X \stackrel{d}{=} Y,$$

it is not true that

$$\frac{X}{Y} \stackrel{d}{=} \frac{Y}{X},$$

even for positive rv's X and Y .

Counterexample 1 (Discrete). Let $\mathbb{P}(X = x, Y = y)$ be given by the following table:

$x \ y$	1	2	3	$f_X(x)$
1	2/36	9/36	1/36	1/3
2	1/36	2/36	9/36	1/3
3	9/36	1/36	2/36	1/3
$f_Y(y)$	1/3	1/3	1/3	1

Obviously $X \stackrel{d}{=} Y$, $X \sim U(\{1, 2, 3\})$, but

$$\mathbb{P}\left(\frac{X}{Y} = 2\right) = \frac{1}{36} \neq \mathbb{P}\left(\frac{Y}{X} = 2\right) = \frac{9}{36}.$$

We construct a similar example for the continuous case:

Counterexample 2 (Continuous). Let U_1, U_2 be i.i.d. and uniformly distributed over the interval $(0, 1)$. Let I be independent of (U_1, U_2) such that $I \sim U(\{0, 1, 2\})$, that is, $\mathbb{P}(I = i) = 1/3$, $i = 0, 1, 2$. Let also $J = I + 1$ if $I = 0$ or $I = 1$, and $J = 0$ if $I = 2$, so that $J \sim U(\{0, 1, 2\})$, i.e.,

$$J \stackrel{d}{=} I.$$

It is clear that (I, J) and (U_1, U_2) are independent random vectors. We now define

$$(X, Y) = (I + U_1, J + U_2).$$

Clearly X and Y have the same distribution, $U(0, 3)$. The joint density of (X, Y) is

$$f(x, y) = \begin{cases} 1/3, & \text{if } x \in (0, 1) \text{ and } y \in (1, 2), \\ 1/3, & \text{if } x \in (1, 2) \text{ and } y \in (2, 3), \\ 1/3, & \text{if } x \in (2, 3) \text{ and } y \in (0, 1), \\ 0, & \text{otherwise.} \end{cases}$$

This density is not exchangeable, $f(x, y) \neq f(y, x)$, and the ratios X/Y and Y/X are not identically distributed, e.g.,

$$\mathbb{P}\left(\frac{X}{Y} \leq 1\right) = \frac{2}{3}, \quad \mathbb{P}\left(\frac{Y}{X} \leq 1\right) = \frac{1}{3}.$$

In both examples given above, the vector (X, Y) fails to be exchangeable. We say that (X, Y) are exchangeable if

$$(X, Y) \stackrel{d}{=} (Y, X). \tag{1}$$

Clearly, if (1) holds then for any (Borel) $g : \mathbb{R}^2 \mapsto \mathbb{R}$ we have

$$g(X, Y) \stackrel{d}{=} g(Y, X),$$

and in particular, for $g(x, y) = x/y$ we get

$$\frac{X}{Y} \stackrel{d}{=} \frac{Y}{X},$$

provided that $\mathbb{P}(X = 0) = 0$. Therefore, we have the following.

Fact 1. If the vector (X, Y) is exchangeable and $\mathbb{P}(X = 0) = 0$ (and thus, $\mathbb{P}(Y = 0) = 0$, by exchangeability), then

$$\frac{X}{Y} \stackrel{d}{=} \frac{Y}{X}. \tag{2}$$

In particular, if X and Y are i.i.d. with $\mathbb{P}(X = 0) = 0$ then (2) holds.

We now pose some questions:

Question 1. If T is an rv with $\mathbb{P}(T = 0) = 0$ such that

$$T \stackrel{d}{=} \frac{1}{T}, \tag{3}$$

is it true that there exist an exchangeable vector (X, Y) with $\mathbb{P}(X = 0) = 0$ such that

$$T \stackrel{d}{=} \frac{X}{Y} ?$$

Question 2 (stronger). If T is an rv with $\mathbb{P}(T = 0) = 0$ such that (3) holds, is it true that there exist i.i.d. rv's X, Y with $\mathbb{P}(X = 0) = 0$ such that

$$T \stackrel{d}{=} \frac{X}{Y} ?$$

The answer to the Question 1 is in the affirmative, in contrast to the Question 2. To see this, consider a symmetric Bernoulli r.v. I with $\mathbb{P}(I = 0) = \mathbb{P}(I = 1) = 1/2$, and define the random vector

$$(X, Y) = (W \cdot T^I, W \cdot T^{1-I}) = (W[(1-I) + IT], W[I + (1-I)T]). \quad (4)$$

In (4), W is **any random variable** with $\mathbb{P}(W = 0) = 0$, for example $W \equiv 1$ or $W \sim N(\mu, \sigma^2)$, and the r.v.'s T, W, I are independent. Since

$$I \stackrel{d}{=} 1 - I$$

it follows that

$$(T, W, I) \stackrel{d}{=} (T, W, 1 - I),$$

and thus, for any Borel $\mathbf{g} = (g_1, g_2) : \mathbb{R}^3 \mapsto \mathbb{R}^2$ we have

$$\mathbf{g}(T, W, I) \stackrel{d}{=} \mathbf{g}(T, W, 1 - I)$$

i.e.,

$$(g_1(T, W, I), g_2(T, W, I)) \stackrel{d}{=} (g_1(T, W, 1 - I), g_2(T, W, 1 - I)). \quad (5)$$

Applying (5) to $\mathbf{g}(t, w, i) = (wt^i, wt^{1-i})$ we get, in view of (4), that

$$(X, Y) \stackrel{d}{=} (Y, X),$$

so that (X, Y) is an exchangeable random vector. Moreover, by definition,

$$\frac{X}{Y} = \frac{T^I}{T^{1-I}}.$$

Finally, for any $t \in \mathbb{R}$ we have

$$\begin{aligned} \mathbb{P}\left(\frac{X}{Y} \leq t\right) &= \frac{1}{2} \mathbb{P}\left(\frac{T^I}{T^{1-I}} \leq t \mid I = 0\right) + \frac{1}{2} \mathbb{P}\left(\frac{T^I}{T^{1-I}} \leq t \mid I = 1\right) \\ &= \frac{1}{2} \mathbb{P}\left(\frac{1}{T} \leq t\right) + \frac{1}{2} \mathbb{P}(T \leq t) \\ &= \mathbb{P}(T \leq t), \end{aligned}$$

because, from (3),

$$\mathbb{P}\left(\frac{1}{T} \leq t\right) = \mathbb{P}(T \leq t).$$

Therefore,

$$\frac{X}{Y} \stackrel{d}{=} T,$$

as it was to be shown.

Note. If the r.v. W is taken to be absolutely continuous then the vector (X, Y) is absolutely continuous. This shows that there always exists an absolutely continuous exchangeable solution even if T is discrete.

We now show that there exists an r.v. T which satisfies (3), but T has not a representation of the form

$$T \stackrel{d}{=} \frac{X}{Y}, \quad \text{with } X, Y \text{ being i.i.d.} \quad (6)$$

Indeed, consider the positive r.v. T for which $\log(T)$ is uniformly distributed over the interval $(-1, 1)$, i.e., $U = \log(T) \sim U(-1, 1)$ — **log-uniform**. Clearly T satisfies (3). If (6) was true then U should have the representation

$$U \stackrel{d}{=} X_1 - X_2, \quad \text{with } X_1 = \log(|X|), \quad X_2 = \log(|Y|). \quad (7)$$

Since X and Y are i.i.d. it follows that X_1 and X_2 in (7) are also i.i.d., but this is impossible. Indeed, if $\phi(t)$ is the common characteristic function of X_1 and X_2 then the characteristic function of $X_1 - X_2$ is, by independence,

$$\phi_{X_1 - X_2}(t) = \phi(t)\phi(-t) = \phi(t)\overline{\phi(t)} = |\phi(t)|^2 \geq 0.$$

On the other hand, the characteristic function of U is

$$\phi_U(t) = \frac{\sin(t)}{t},$$

which assumes both positive and negative values.