

$$X \sim N(\mu, \sigma^2)$$

$$f_X(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)$$

$$\mu \in \mathbb{R}, \sigma^2 > 0$$

$$-\infty < x < +\infty$$

$$\sigma = \sqrt{\sigma^2}, \quad \text{Υπολογισμοί:}$$

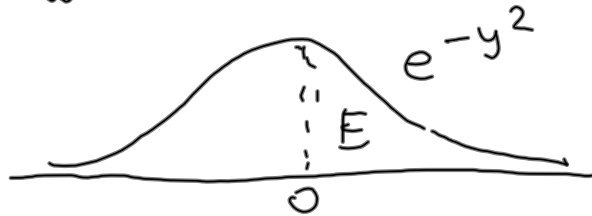
$$\int_{-\infty}^{\infty} \exp\left(-\left(\frac{x-\mu}{\sigma\sqrt{2}}\right)^2\right) dx$$

$$\frac{x-\mu}{\sigma\sqrt{2}} = y$$

$$= \underbrace{\sigma\sqrt{2}} \int_{-\infty}^{\infty} e^{-y^2} dy = \sigma\sqrt{2\pi}$$

$$x = \mu + y\sigma\sqrt{2}$$
$$dx = \sigma\sqrt{2} dy$$

$$I = \int_{-\infty}^{\infty} e^{-y^2} dy = 2E = 2 \int_0^{\infty} e^{-y^2} dy$$



$$\underline{y^2 = t} \quad y = \sqrt{t} \quad dy = \frac{1}{2\sqrt{t}} dt \quad t \in (0, \infty)$$

$$I = 2 \cdot \int_0^{\infty} \frac{1}{2\sqrt{t}} e^{-t} dt = \int_0^{\infty} t^{\frac{1}{2}-1} e^{-t} dt$$

$$\Gamma(\alpha) = \int_0^{\infty} t^{\alpha-1} e^{-t} dt \quad = \Gamma\left(\frac{1}{2}\right) = \sqrt{\pi}$$

Πρόταση: Αν $X \sim N(\mu, \sigma^2)$ τότε

$$\eta \quad Z \stackrel{\text{op}}{=} \frac{X - \mu}{\sigma} \sim N(0, 1).$$

Αντίστροφα, αν $η \quad Z \sim N(0, 1)$ τότε $η$

$$X \stackrel{\text{op}}{=} \mu + \sigma Z \sim N(\mu, \sigma^2)$$

Γενικότερα: Αν $X \sim N(\mu, \sigma^2)$, και $\alpha \in \mathbb{R}^+$,

$$\beta \in \mathbb{R} \Rightarrow Y = \alpha X + \beta \sim N(\alpha\mu + \beta, \alpha^2\sigma^2).$$

$$X \sim N(\mu, \sigma^2), \quad \text{Θ ζρω} \quad Z = \frac{X - \mu}{\sigma}$$

$$P(Z \leq t) = F_Z(t) = P\left(\frac{X - \mu}{\sigma} \leq t\right)$$

$$\stackrel{\sigma > 0}{=} P(X \leq \mu + \sigma t) = F_X(\mu + \sigma t)$$

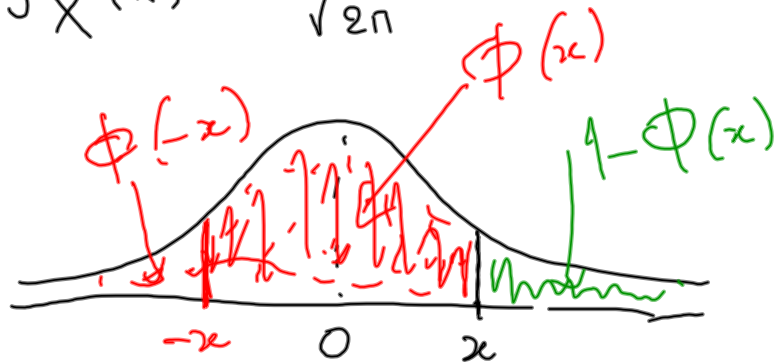
$$f_Z(t) = F_Z'(t) = F_X'(\mu + \sigma t) \cdot (\mu + \sigma t)'$$

$$= \sigma f_X(\mu + \sigma t) = \frac{\sigma}{\sigma \sqrt{2\pi}} \exp\left(-\frac{(\mu + \sigma t - \mu)^2}{2\sigma^2}\right)$$

$$= \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{t^2}{2}\right) \quad Z \sim N(0, 1). \quad \square$$

Τύπος. Κανονική $\equiv N(0,1)$

$$\varphi(x) = f_X(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2} \quad x \in \mathbb{R}$$



$$\Phi(x) \stackrel{\text{op}}{=} \int_{-\infty}^x \varphi(t) dt \quad \varphi = \Phi'$$
$$\Phi(0) = \frac{1}{2}$$

$$\Phi(-x) = 1 - \Phi(x) \quad \forall x \in \mathbb{R}$$

για $x=0$ παίρνω $\Phi(0) = 1/2$

Μέση τιμή: $Z \sim N(0,1)$ τότε

(i) $E(Z) = 0$ (ii) $\text{Var}(Z) = 1$.

$$X = \mu + \sigma Z \quad (\text{όταν } X \sim N(\mu, \sigma^2))$$

$$E(X) = \mu + \sigma \overset{0}{E(Z)} = \mu, \quad \text{Var}(X) = \text{Var}(\mu + \sigma Z) \\ = \sigma^2 \text{Var}(Z) = \sigma^2.$$

$$Z \sim N(0,1)$$

$$E(\hat{Z}) = \int_{-\infty}^{\infty} \underbrace{x \varphi(x)}_{\text{νεφερώ}} dx = 0$$

$$E(\hat{Z}^2) = \int_{-\infty}^{\infty} x^2 \varphi(x) dx = 2 \int_0^{\infty} x^2 \varphi(x) dx$$

$$= \frac{2}{\sqrt{2\pi}} \int_0^{\infty} x^2 e^{-x^2/2} dx$$

$$\begin{aligned} x^2/2 &= y \\ x &= \sqrt{2y} \end{aligned}$$

$$= \frac{2}{\sqrt{2\pi}} \int_0^{\infty} 2y e^{-y} \frac{1}{\sqrt{2y}} dy$$

$$dx = \frac{1}{2\sqrt{2y}} \cdot 2 dy$$

$$= \frac{2}{\sqrt{\pi}} \int_0^{\infty} \sqrt{y} e^{-y} dy = \frac{2}{\sqrt{\pi}} \cdot \int_0^{\infty} y^{\frac{3}{2}-1} e^{-y} dy$$

$$\Gamma(\alpha+1) = \alpha \Gamma(\alpha)$$

$$= \frac{2}{\sqrt{\pi}} \Gamma\left(\frac{1}{2}+1\right) = \frac{2}{\sqrt{\pi}} \cdot \frac{1}{2} \cdot \frac{\Gamma\left(\frac{1}{2}\right)}{\sqrt{\pi}} = 1.$$

$$\text{Var}(Z) = E(Z^2) - (E(Z))^2 = 1 - 0^2 = 1.$$

Π. 8.

$$\Phi(1) = 0.8413 \quad \Phi(1.2) = 0.8849$$

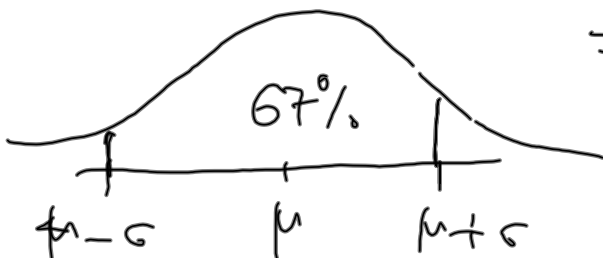
$$\Phi(3) = 0.9997$$

$$P(\mu - \sigma < X < \mu + \sigma) = ? \quad \underline{X \sim N(\mu, \sigma^2)}$$

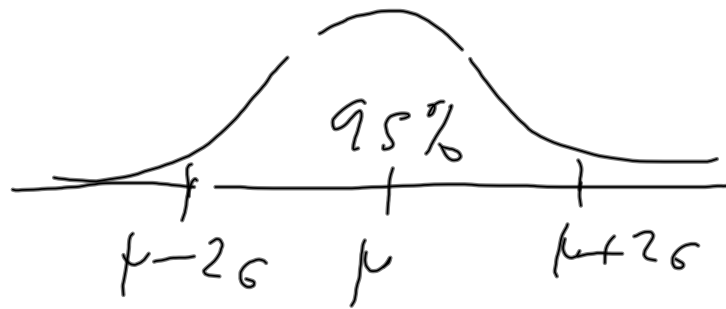
$$\begin{aligned} &\parallel \\ &P\left(-1 < \frac{X - \mu}{\sigma} < 1\right) = \Phi(1) - \Phi(-1) \\ &\quad \parallel \\ &\quad Z \sim N(0, 1) \end{aligned}$$

$$\begin{aligned} &X_{GW} \\ &\parallel \\ &= \Phi(1) - \Phi(-1) = \Phi(1) - (1 - \Phi(1)) \end{aligned}$$

$$= 2\Phi(1) - 1 \approx 67\%$$



$$\begin{aligned}\Phi(\mu - 2\sigma < X < \mu + 2\sigma) &= 2\Phi(2) - 1 \\ &= 2 \cdot 0.9773 - 1 \approx 95\%\end{aligned}$$



Άσκησης κεφ 4

4.1, 4.2, 4.5, 4.7, 4.8, 4.11, 4.12,

4.13, 4.14, 4.15, 4.17, 4.18

4.7] $X \sim \text{Exp}(\theta)$ ($\theta > 0$). Βρείτε
την συνκνότητα της $Y = 1 - e^{-\theta X}$.

$$P(Y \leq y) = P(1 - e^{-\theta X} \leq y) \quad y \in (0, 1)$$

$$= P(e^{-\theta X} \geq 1 - y)$$

$$= P(-\theta X \geq \log(1 - y))$$

$$= P(X \leq -\frac{1}{\theta} \log(1 - y)) = F_X\left(-\frac{1}{\theta} \log(1 - y)\right)$$

$$\begin{aligned} \hat{f}_Y(y) &= F'_Y(y) = f_X\left(-\frac{1}{\theta} \log(1 - y)\right) \cdot \left(-\frac{1}{\theta} \log(1 - y)\right)' \\ &= \theta e^{-\theta \cdot \left(-\frac{1}{\theta} \log(1 - y)\right)} \cdot \frac{-1}{\theta} \cdot \frac{1}{1 - y} \cdot (1 - y)' = \frac{1 - y}{1 - y} = 1. \end{aligned}$$

$$f_Y(y) = 1, \quad \text{για } y \in (0,1)$$

$$(\text{ο } \alpha \text{ λ } \text{ο } \beta). \quad Y \sim U(0,1)$$

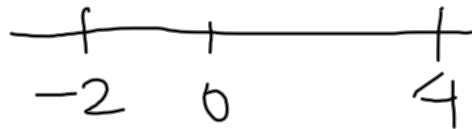
$$4.1) \quad \underline{X \sim U(\alpha, \beta)}, \quad E(X) = 1, \quad \text{Var}(X) = 3$$

$$\alpha, \beta = ?$$

$$\left. \begin{array}{l} 1 = \mu = \frac{\alpha + \beta}{2} \\ 3 = \sigma^2 = \frac{(\beta - \alpha)^2}{12} \end{array} \right\} \begin{array}{l} \alpha + \beta = 2 \\ (\beta - \alpha)^2 = 36 \\ \alpha < \beta \end{array} \right\} \begin{array}{l} \alpha + \beta = 2 \\ \alpha = \beta \\ \beta = 4 \\ \alpha = -2 \end{array}$$

$$X \sim U(-2, 4)$$

$Y = |X|$. Βρείτε την συν. της Y .



$Y = |X| \in (0, 4)$, $y \in (0, 4)$

$$F_Y(y) = P(|X| \leq y) = P(-y \leq X \leq y)$$

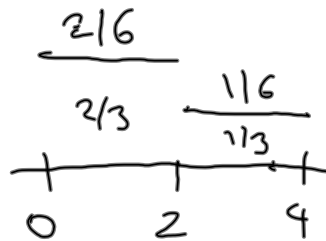
$$= F_X(y) - F_X(-y-) \stackrel{\text{σω.}}{=} F_X(y) - F_X(-y)$$

$$f_Y(y) = (F_X(y) - F_X(-y))'$$

$$= f_X(y) + f_X(-y)$$

$$f_X(x) = \begin{cases} 1/6, & x \in (-2, 4) \\ 0, & \text{σίστησφ.} \end{cases}$$

$$= \begin{cases} 1/6 + 1/6, & y \in (0, 2) \\ 1/6 + 0, & y \in (2, 4) \end{cases}$$



Βρείτε $E(Y)$, $\text{Var}(Y)$

$$Y = |X|$$

$$\begin{aligned} E(Y) &= \int_0^4 y f_Y(y) dy = \int_0^2 \frac{2}{6} y + \int_2^4 \frac{1}{6} \cdot y \\ &= \dots = 5/3 \end{aligned}$$

$$E(Y^2) = \int_0^2 y^2 \frac{2}{6} + \int_2^4 y^2 \frac{1}{6} = 4 = E X^2$$

$$\text{Var}(Y) = E(Y^2) - (E(Y))^2 = \frac{11}{9}.$$