

Ειδικές Συνεχείς Κατανομές

$$F(x) = \int_{-\infty}^x f(t) dt \quad \forall x$$

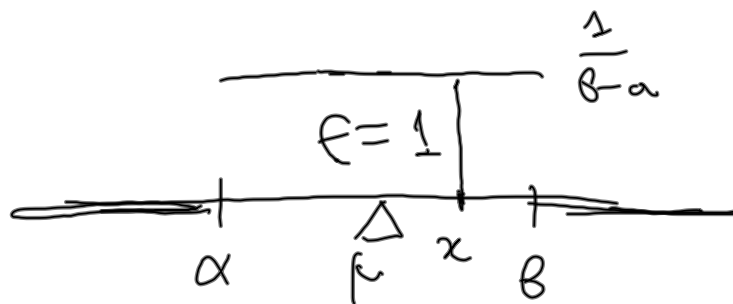
f : πυκνότητα (πιθανότητας)

$$F' = f, \quad \int_{-\infty}^{\infty} f = 1 \quad (f \geq 0)$$

Ομοιόμορπη στο (α, β) $-\infty < \alpha < \beta < \infty$

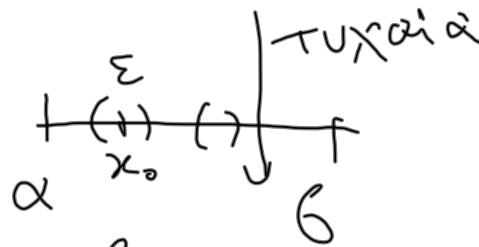
$$X \sim \mathcal{U}(\alpha, \beta)$$

$$\Leftrightarrow f_X(x) = \begin{cases} \frac{1}{\beta - \alpha}, & \alpha < x < \beta \\ 0, & \text{αλλιώς} \end{cases}$$



$$F(x) = \begin{cases} 0, & x \leq a \\ \frac{x-a}{b-a}, & a \leq x \leq b \\ 1, & x \geq b \end{cases}$$

$$X \sim U(a, b)$$



$$\mu = \int_{-\infty}^{\infty} x f(x) dx = \int_a^b x \cdot \frac{1}{b-a} dx$$

$$= \frac{1}{\beta - \alpha} \left. \frac{x^2}{2} \right|_{\alpha}^{\beta} = \frac{\beta^2 - \alpha^2}{2(\beta - \alpha)} = \frac{\alpha + \beta}{2}$$

$$\mu = \frac{\alpha + \beta}{2}$$

$$E(X^n) = \int_{-\infty}^{\infty} x^n f(x) dx$$

$$= \frac{1}{\beta - \alpha} \int_{\alpha}^{\beta} x^n dx = \frac{\beta^{n+1} - \alpha^{n+1}}{(n+1)(\beta - \alpha)}$$

$$E(X^2) = \frac{\beta^3 - \alpha^3}{3(\beta - \alpha)} = \frac{\alpha^2 + \alpha\beta + \beta^2}{3}$$

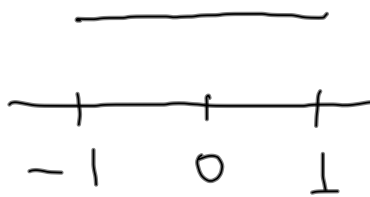
$$\sigma^2 = \text{Var}(X) = E(X^2) - \mu^2$$

$$= \frac{\alpha^2 + \alpha\beta + \beta^2}{3} - \left(\frac{\alpha + \beta}{2}\right)^2$$

$$= \dots = \frac{(\beta - \alpha)^2}{12} = \sigma^2 \quad \mu = \frac{\alpha + \beta}{2}$$

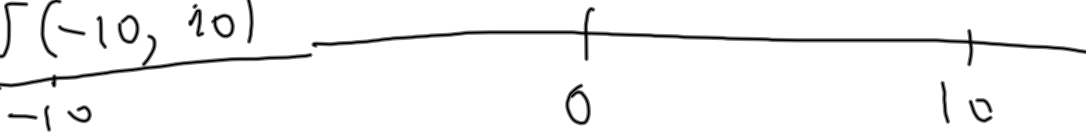
$$\sigma = \frac{\beta - \alpha}{\sqrt{12}}$$

$\mathcal{N}(-1, 1)$



$$\mu = 0, \quad \sigma^2 = \frac{1}{3}$$

$\mathcal{N}(-10, 10)$



$$\sigma^2 = \frac{(10 - (-10))^2}{12} = \frac{2^2 \cdot 10^2}{12} = \frac{10^2}{3} = 100 \cdot \frac{1}{3}$$

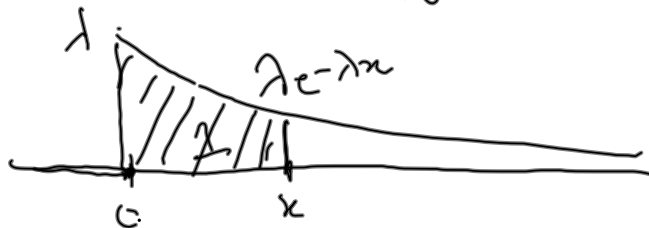
Συθετική με παράμετρο $\lambda > 0$

$X \sim \text{Exp}(\lambda)$ είναι η πυκνότητα

$$f_X(x) = \begin{cases} \lambda e^{-\lambda x}, & x > 0 \\ 0, & \text{διαφ.} \end{cases}$$

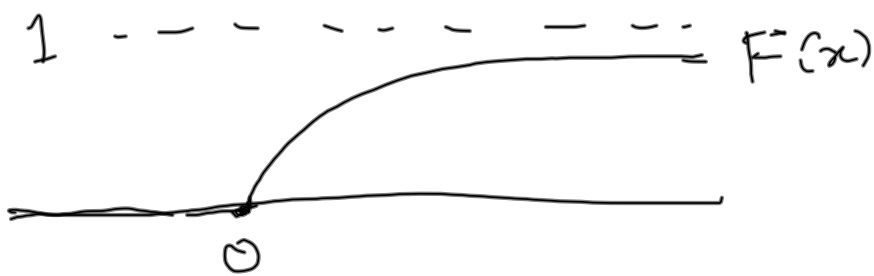
$$f \geq 0 \quad \int_{-\infty}^{\infty} f(x) dx = \int_0^{\infty} \lambda e^{-\lambda x} dx$$

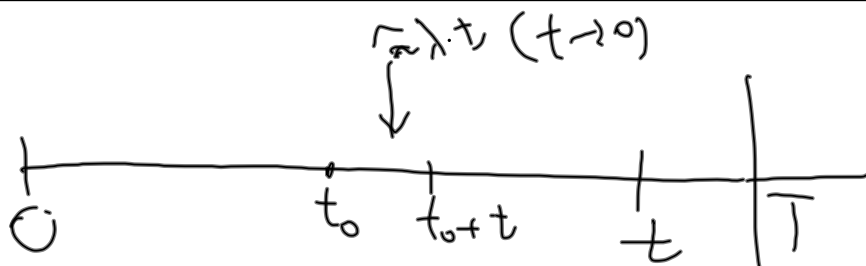
$$= \int_0^{\infty} (-e^{-\lambda x})' dx = -e^{-\lambda x} \Big|_0^{\infty} = -0 + e^0 = 1$$



$$F(x) = \int_0^x (-e^{-\lambda t})' dt = -e^{-\lambda x} + 1$$

$$F(x) = \begin{cases} 0, & x \leq 0 \\ 1 - e^{-\lambda x}, & x \geq 0 \end{cases}$$





$X(t) = \#$ συμβάντων στο $[0, t]$

$$X(t) \sim \text{Poisson}(\lambda \cdot t)$$

$$P(X(t) = x) = e^{-\lambda t} \frac{(\lambda t)^x}{x!}, \quad x = 0, 1, \dots$$

$$P(X(t) = 0) = e^{-\lambda t} = 1 - (1 - e^{-\lambda t}) = 1 - F_X(t)$$

$$\{X(t)=0\} = \{T > t\}$$

$T =$ χρονική στιγμή του συμβάντος.

$$P(T > t) = P(X(t)=0) = e^{-\lambda t}$$

$$P(T \leq x) = 1 - e^{-\lambda x} = F_T(x)$$

δηλ. $T \sim \text{Exp}(\lambda)$

$$x, y > 0$$

$$P(X > x+y | X > x) = P(A|B) \cdot$$

$$A = \{\omega: X(\omega) > x+y\}, \subseteq B = \{\omega: X(\omega) > x\}$$

$$\cdot = \frac{P(A \cap B)}{P(B)} = \frac{P(X > x+y)}{P(X > x)} = \frac{\exp(-\lambda) e^{-\lambda(x+y)}}{e^{-\lambda x}}$$

$$X \sim \text{Exp}(\lambda) \quad (P(X > t) = e^{-\lambda t})$$

$$= e^{-\lambda y}$$

$$= P(X > y)$$

$$P(X > x+y | X > x) = \underline{P(X > y)}$$

$$\forall x, y > 0.$$

$$\mu = \frac{1}{\lambda}, \quad \sigma^2 = \frac{1}{\lambda^2}$$

Μέση τιμή /
διασπορά
εξαρτημένης.

$$\mu = \int_0^{\infty} x \lambda e^{-\lambda x} dx = \int_0^{\infty} x (-e^{-\lambda x})' dx$$

$$= x (-e^{-\lambda x}) \Big|_0^{\infty} - \int_0^{\infty} (x)' (-e^{-\lambda x}) dx$$

$$= \int_0^{\infty} e^{-\lambda x} dx = \frac{1}{\lambda}$$

$$E(X^2) = \int_0^{\infty} x^2 \lambda e^{-\lambda x} dx = \int_0^{\infty} x^2 (-e^{-\lambda x})' dx$$

$$= \dots = \frac{2}{\lambda^2} \quad \sigma^2 = \frac{2}{\lambda^2} - \left(\frac{1}{\lambda}\right)^2 = \frac{1}{\lambda^2}$$

$$E(X^n) = \frac{n!}{\lambda^n} \quad n=1,2,\dots$$

$\Gamma(\alpha, \lambda)$

Σωάρτηση $\Gamma(\alpha)$ του Euler

$$\forall \alpha > 0 \quad \Gamma(\alpha) \stackrel{\text{op.}}{=} \int_0^{\infty} w^{\alpha-1} e^{-w} dw,$$

$$\Gamma(1) = \int_0^{\infty} w^{1-1} e^{-w} dw = \int_0^{\infty} e^{-w} dw = 1$$

$$\Gamma(\alpha+1) = \int_0^{\infty} w^{\alpha} (-e^{-w})' dw$$

$$= -\cancel{e^{-w} w^{\alpha}} \Big|_0^{\infty} + \int_0^{\infty} \alpha w^{\alpha-1} e^{-w} dw$$

$\Gamma(\alpha+1) = \alpha \Gamma(\alpha)$	$\Gamma(1) = 1$
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$\alpha > 0$

$$\Gamma(2) = 1 \cdot \Gamma(1) = 1 = 1!$$

$$\Gamma(3) = 2 \cdot \Gamma(2) = 2 \cdot 1 = 2 = 2!$$

$$\Gamma(4) = 3 \Gamma(3) = 3 \cdot 2 = 6 = 3!$$

$$\Gamma(5) = 4 \Gamma(4) = 4 \cdot 6 = 24 = 4!$$

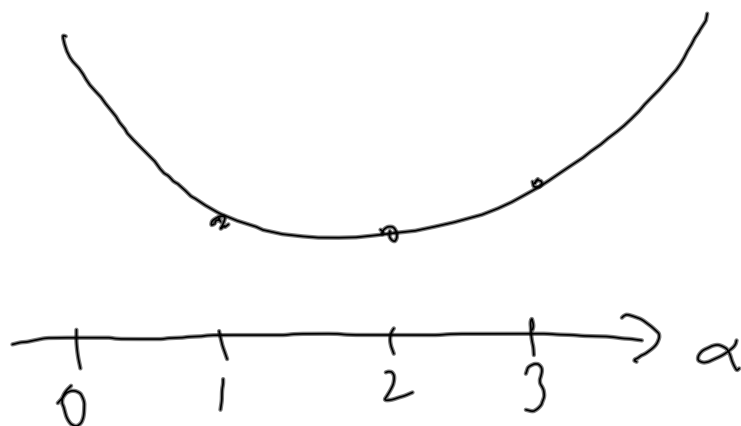
$$\Gamma(6) = 5 \Gamma(5) = 5 \cdot 24 = 120 = 5!$$

$$\Gamma(n) = (n-1)! \quad n=1, 2, \dots$$

$$\Gamma\left(\frac{1}{2}\right) = \sqrt{\pi} \quad (\text{Euler})$$

$$\Gamma\left(\frac{3}{2}\right) = \Gamma\left(1 + \frac{1}{2}\right) = \frac{1}{2} \Gamma\left(\frac{1}{2}\right) = \frac{\sqrt{\pi}}{2}$$

$$\Gamma\left(\frac{5}{2}\right) = \frac{3}{2} \Gamma\left(\frac{3}{2}\right) = \frac{3}{2} \frac{\sqrt{\pi}}{2} = \frac{3}{4} \sqrt{\pi}$$



$$X \sim \Gamma(\alpha, \lambda) \quad \text{ότι α ν}$$

$$f_X(x) = \frac{\lambda^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-\lambda x}, \quad x > 0$$

$$\int_0^\infty x^{\alpha-1} e^{-\lambda x} dx = \int_0^\infty \left(\frac{w}{\lambda}\right)^{\alpha-1} e^{-w} \frac{1}{\lambda} dw$$

$\lambda x = w \quad x = \frac{w}{\lambda} \quad dx = \frac{1}{\lambda} dw$

$$= \frac{1}{\lambda^\alpha} \underbrace{\int_0^\infty w^{\alpha-1} e^{-w} dw}_{\Gamma(\alpha)} = \frac{\Gamma(\alpha)}{\lambda^\alpha}$$

Ειδικά όταν $\alpha = n \in \{1, 2, \dots\}$

$$f_X(x) = \frac{\lambda^n}{(n-1)!} x^{n-1} e^{-\lambda x}, \quad x > 0$$

$$\text{Erlang}(n, \lambda) \equiv \Gamma(n, \lambda)$$

$$n=1, \quad f(x) = \lambda e^{-\lambda x}, \quad x > 0$$

$$\text{Exp}(\lambda) \equiv \Gamma(1, \lambda)$$

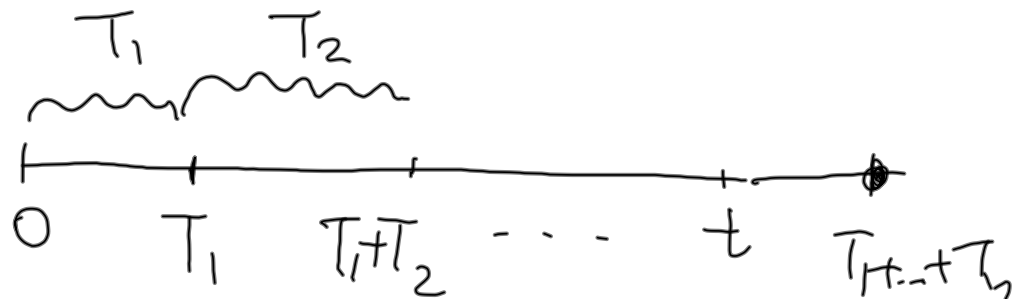
$$\mu = \frac{\alpha}{\lambda}, \quad \sigma^2 = \frac{\alpha}{\lambda^2}$$

$$\Gamma(\alpha, \lambda)$$

$$\mu = \frac{n}{\lambda}, \quad \sigma^2 = \frac{n}{\lambda^2}$$

$$\text{Exp}(n, \lambda)$$

$$\{T_1 + \dots + T_n > t\} \Leftrightarrow \{X(t) \leq n-1\}$$



$$T_1, \dots, T_n \sim \text{Exp}(\lambda) \Rightarrow T_1 + \dots + T_n \sim \text{Erd}(\eta, \lambda)$$

$$P(X(t) \leq n-1) \\ = e^{-\lambda t} \sum_{k=0}^{n-1} \frac{(\lambda t)^k}{k!}$$

$$= P(T_1 + \dots + T_n > t) = 1 - P(\dots \leq t)$$

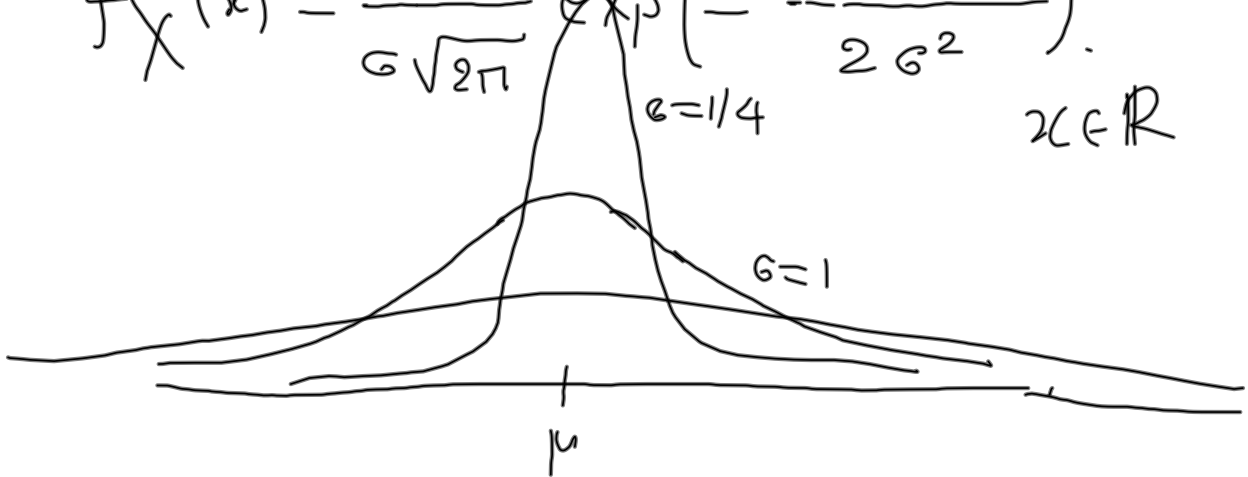
$$= 1 - \int_0^t \underbrace{\frac{\lambda^n}{(n-1)!} x^{n-1} e^{-\lambda x}}_{\text{pdf of } T_n} dx$$

$$F_X(x) = 1 - e^{-\lambda x} \sum_{k=0}^{n-1} \frac{(\lambda x)^k}{k!}$$

όραση $X \sim \text{Erlang}(n, \lambda)$.

$N(\mu, \sigma^2)$ Κανονική μ, σ^2 .
 $\mu \in \mathbb{R}, \sigma^2 > 0$

$$f_X(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right), \quad x \in \mathbb{R}$$

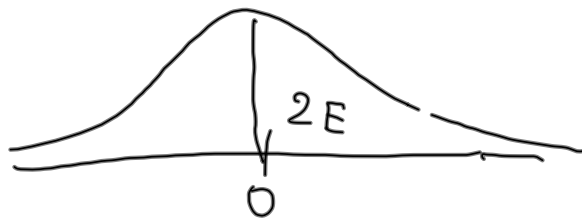


$$I = \int_{-\infty}^{\infty} \exp\left(-\left(\frac{x-\mu}{\sigma\sqrt{2}}\right)^2\right) dx$$

$$y = \frac{x-\mu}{\sigma\sqrt{2}} \Rightarrow x = \mu + \sigma\sqrt{2} \cdot y$$
$$dx = \sigma\sqrt{2} \cdot dy$$

$$= \int_{-\infty}^{\infty} e^{-y^2} \sigma\sqrt{2} dy = \sigma\sqrt{2} \cdot \underbrace{\int_{-\infty}^{\infty} e^{-y^2} dy}_{\sqrt{\pi}}$$

$$\int_{-\infty}^{\infty} e^{-y^2} dy = 2 \int_0^{\infty} e^{-y^2} dy \quad \text{ευθεία τμήτ.$$



$$y^2 = w$$

$$y = \sqrt{w}$$

$$dy = \frac{1}{2\sqrt{w}} dw$$

$$= 2 \int_0^{\infty} e^{-w} \frac{1}{2\sqrt{w}} dw = \int_0^{\infty} w^{\frac{1}{2}-1} e^{-w} dw = \Gamma\left(\frac{1}{2}\right)$$

$$N(0,1) \quad \mu=0, \quad \sigma^2=1$$

Τυποποίηση της κανονικής

$$f_X(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}, \quad x \in \mathbb{R}$$

$$= \varphi(x)$$

$$\Phi(x) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-t^2/2} dt$$
